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Abbreviations:

AADT = annual average daily traffic

Cd = cadmium

CTS = California Teachers Study

FFQ = food frequency questionnaire

GIS = geographic information system

GM = geometric mean

ICC = intraclass correlation coefficient

LOD = limit of detection

 R^2 = coefficient of determination

U-Cd = urinary cadmium

VKT = vehicle kilometers traveled

ABSTRACT

Background: Cadmium is a toxic metal associated with increased morbidity and mortality.

Urinary cadmium concentration is considered a biomarker of long-term exposure.

Objectives: Our objectives were to evaluate the within-person correlation among repeat samples and identify predictors of urinary cadmium concentrations.

Methods: Urinary cadmium concentrations (μg/L) were measured in 24-hour urine samples collected from 296 women enrolled in the California Teachers Study in 2000 and a second 24-hour sample collected 3 – 9 months later from 141 of the participants. Lifestyle and sociodemographic characteristics were obtained via questionnaires. The Total Diet Study database was used to quantify dietary cadmium intake based on a food frequency questionnaire. We estimated environmental cadmium emissions near participants' residences using a geographic information system.

Results: The geometric mean urinary cadmium concentration was $0.27\mu g/L$ and the range was $0.1-3.6 \mu g/L$. The intraclass correlation among repeat samples from an individual was 0.50. The use of a single 24-hour urine specimen to characterize Cd exposure in a case-control study would result in an observed odds ratio of 1.4 for a true odds ratio of 2.0. Urinary cadmium concentration increased with creatinine, age and lifetime pack-years of smoking among ever smokers or lifetime intensity-years of passive smoking among nonsmokers, while it decreased with greater alcohol consumption and number of previous pregnancies. These factors explained 42-44% of the variability in urinary cadmium concentrations.

Conclusion: Urinary cadmium levels varied with several individual characteristics and a single measurement of urinary cadmium in a 24-hour sample did not accurately reflect medium to long-term body burden.

BACKGROUND

Cadmium (Cd) is a toxic, bioaccumulating, and somewhat persistent metal released into the environment during mining operations, industrial processes, and as a byproduct of oil combustion (ATSDR 1999). Non-occupational Cd exposure, assessed using urinary Cd levels from a single sample, has been associated with kidney disease (Jarup and Akesson 2009; Suwazono et al. 2006), cardiovascular disease (Peters et al. 2010), dental caries (Arora et al. 2008), decreases in bone mineral density (Satarug and Moore 2004), and increased mortality (Menke et al. 2009; Nawrot et al. 2008). Cd is classified as a human lung carcinogen (IARC 1993), and has been associated with increased overall cancer mortality (Adams et al. 2012) and the incidence of breast (Gallagher et al. 2010; Julin et al. 2012; McElroy et al. 2006) and endometrial cancers (Akesson et al. 2008).

Cd is stored in the liver and kidneys, has a biological half-life of 10–30 years, and is absorbed via inhalation and ingestion (ATSDR 1999). Absorption of Cd in the gastrointestinal tract is limited (3-10%) while absorption from the deep areas of the lung is high (50-90%), suggesting that inhalation may be an important route of exposure (Waalkes 2003). The principal source of exposure for smokers in non-industrial settings is inhalation of cigarette smoke (CDC 2005) as smoking may double the daily intake of Cd compared with not smoking (ATSDR 1999). For nonsmokers, the principal source of Cd is ingestion of contaminated plant-based foods (CDC 2005). Women generally have higher internal Cd levels than men because depleted iron stores and iron deficiency, common among women of childbearing age, increases the intestinal uptake of Cd (Vahter et al. 2002).

Numerous studies have cited urinary Cd (U-Cd) concentration as a reliable measure of cumulative lifetime exposure (Julin et al. 2011; McElroy et al. 2006; Nawrot et al. 2006). Accurate assessment of long-term Cd exposure is important because U-Cd levels have been associated with health outcomes thought to have a long latency period such as cancer and cardiovascular disease. However, short-term Cd exposure levels could also be important for studies of prenatal exposure and developmental effects in children. U-Cd concentrations have been shown to be correlated with age (Hellstrom et al. 2004), gender (Hellstrom et al. 2004), iron deficiency (Berglund et al. 1994), parity (Akesson et al. 2002), smoking status (Hellstrom et al. 2004; Ikeda et al. 2005), second-hand smoke (Willers et al. 2005), and dietary intake of Cd (Adams et al. 2011; Choudhury et al. 2001; Julin et al. 2011; Shimbo et al. 2000). However, most of these exposure studies have relied on single spot urine samples instead of repeated 24hour urine collections. In addition, sources of Cd exposure among non-occupationally exposed and mostly nonsmoking women have not been well characterized. The objective of this study was to identify determinants of urinary Cd using repeat 24-hour urine collections and exposure information from a variety of sources including self-reports, environmental databases, and a dietary contaminant database in a sample of women enrolled in the largely non-smoking California Teachers Study (CTS) cohort.

METHODS

Study population and questionnaire data

Our study population consisted of 296 women participating in a measurement sub-study of the CTS cohort. The cohort includes 133,479 women who were active or retired public school teachers or administrators in 1995 (Bernstein et al. 2002). The sub-study, conducted in 2000,

Alameda, Santa Clara, San Mateo, Santa Cruz, Monterey, or northern San Benito counties in California) and were aged 85 years or younger at baseline in 1995-96 (Gunier et al. 2006; Horn-Ross et al. 2008). Of the 484 women invited to participate in the sub-study, 328 (68%) agreed, 138 refused and 18 were not interviewed for other reasons. All participants provided written informed consent and this research was approved by the Institutional Review Board of the Cancer Prevention Institute of California. Of the 328 participants, 304 (93%) provided a 24-hour urine specimen in 2000. Our analysis is based on 296 of these women with adequate urine volume available for cadmium analysis. Of the 157 women asked, 141 (90%) provided a second 24-hour urine specimen; these samples were collected 3, 6, or 9 months following the initial sample. Both the original and repeat specimens were analyzed for U-Cd.

Information on parity, duration of breastfeeding, and active smoking history was obtained from self-administered questionnaires completed when the CTS cohort was established in 1995-96, and information on the source, setting (household, workplace and social), timing, and dose of passive smoking exposures was obtained from a second survey mailed to CTS participants in 1997. For nonsmokers, we used a measure of lifetime intensity-years of passive smoking calculated by multiplying a qualitative description of smoke intensity (1=a little smoky, 2=fairly smoky or 3=very smoky) by duration of exposure in years (Reynolds et al. 2009). Current age, usual diet and alcohol consumption during the past year, weight, height and current residential address were obtained from sub-study participants at the time of urine sample collection in 2000.

Sample collection and laboratory analysis

Each sub-study participant received a collection kit and was instructed to collect all urine produced in the 24-hour period starting immediately following the in-person interview. The

samples were collected and stored at -20°C up to two weeks until they were thawed, aliquotted and frozen at -70°C. Approximately 9 years elapsed between the sample collection in 2000 and the analysis of Cd concentrations.

Urinary cadmium concentrations (μg/L) were measured using inductively-coupled plasma/ mass spectrometry at a certified commercial laboratory (Pacific Toxicology Laboratories; Chatsworth, CA). The limit of detection (LOD) for U-Cd was 0.1 μg/L. Low and high Cd control standards were included in each batch to evaluate assay accuracy and precision. The within-batch coefficient of variation was <10% and between-batch coefficient was <15%. Creatinine concentrations (g/L) were measured using a modified-rate Jaffe method and were highly correlated with creatinine concentrations measured in 2000 at another laboratory (intraclass correlation coefficient=0.88).

Environmental and dietary exposure assessment

We estimated potential environmental exposure to Cd at the sub-study participants' geocoded residences in 2000 using a geographic information system (GIS) and three available databases for industrial emissions, ambient air concentrations, and vehicle traffic. To estimate exposure to Cd emissions from industrial and commercial facilities, we used 1995 data from the California Air Toxics Emissions Data System which provides latitude/longitude coordinates and annual emissions in pounds self-reported by each facility (CARB 1998). We estimated the distance between a residence and all facilities within five kilometers with reported Cd emissions. Geocoded residences were also linked by census tract to estimated Cd concentrations in ambient air in 1999 from the National Air Toxics Assessment (USEPA 2006). These concentrations were derived using an atmospheric dispersion model that combined emissions inventories with local meteorology (Rosenbaum et al. 1999).

To estimate potential exposure to Cd from vehicle emissions, we obtained traffic count data for 2000 from the California Highway Performance and Monitoring System (CDOT 2007). These data provide the annual average daily traffic (AADT), the average number of vehicles per day traveling in both directions on major roads. For each participant's residence, we calculated traffic density by summing the vehicle kilometers of travel (VKT) within a 300 meter radius buffer by multiplying the AADT by the length of the road segment for each road segment with AADT values within the buffer, then dividing by the buffered area (0.28 km²) to obtain VKT per day per square kilometer (Gunier et al. 2003). We used a 300 meter radius because this approximates the distance at which particulate pollutant concentrations approach background levels (Zhou and Levy 2007).

Dietary Cd intake was assessed via an early version of the 103-item Block95 food frequency questionnaire (FFQ) administered to sub-study participants in 2000 (Block et al. 1986; Block et al. 1990; Horn-Ross et al. 2008). For each food item, frequency of consumption (categories ranging from never to once/day or 5+/day depending on the item) and usual portion size (small, medium, or large relative to a given standard medium portion) were assessed for the previous year (i.e., 1999). The FFQs were self-administered and checked by study staff for completeness. FFQ items were assigned Cd values based on the Total Diet Study market basket surveys conducted between 1991 and 2004 (USFDA 2006). Dietary Cd was not estimated for 3 participants who did not complete the food frequency questionnaire and 6 participants whose reported food consumption was judged to be implausibly low (<600 calories/day) or high (>5,000 calories/day).

Statistical analysis

For seven samples with U-Cd concentration below the limit of detection (LOD), we assigned a concentration equivalent to the LOD $\div \sqrt{2}$ (specifically, 0.07 µg/L). We calculated the creatinineadjusted U-Cd levels (µg/g-Cr) by dividing the U-Cd concentrations (µg/L) by the creatinine concentrations (g/L). We multiplied U-Cd concentration (µg/L) by the total volume of urine collected during the 24-hour period (L/day) to estimate daily Cd output (µg/day). Potential explanatory variables for the variation in U-Cd concentrations included age at the time of urine sample collection (2000) rescaled so that the youngest person had an age of zero years, body mass index (kg/m²; a measure of weight independent of height) and body surface area [(weight(kg) $^{0.425}$ × height(cm) $^{0.275}$ × 0.007184); a measure of body size reflecting muscle mass] (Ruggieri and Rocca 2010) using height and weight from 2000, parity (i.e., still and live births were combined to obtain the total number of full-term pregnancies), total duration of breastfeeding (months), oral contraceptive use (ever/never), and hormone replacement therapy (ever/never) as of 1995-96; lifetime active and passive smoking history (through 1997); usual alcohol consumption (g/d), dietary Cd intake (µg/d), and environmental indicators of potential exposure from traffic, industrial, and commercial sources as of 1999. Because the distribution of U-Cd was skewed, we used the non-parametric Kruskal-Wallis test to make categorical comparisons of the U-Cd distribution from the first urine sample collected from each participant (n=296) and demographic, dietary and environmental characteristics. We tested continuous and ordinal predictor variables for a linear trend in regression models using natural-log transformed U-Cd concentrations as the dependent variable to normalize the U-Cd distribution and adjusting for age at urine collection since age was a strong predictor of U-Cd levels.

Variance components models with random intercepts for each participant were used to determine the intraclass correlation coefficient (ICC) of U-Cd concentrations from repeated samples collected from the same individual. We calculated the ratio of the within- and between-person variance (λ =variance ratio) and the attenuation bias that would result from measurement error in a case-control study using a single measure of U-Cd to estimate exposure where the expected value of the logistic regression coefficient for U-Cd (β_E) is related to the true regression coefficient (β_T), the variance ratio (λ) and the number of measurements per person (n) by the equation $\beta_E = \beta_T/(1+\lambda/n)$ and the attenuation bias in a case-control study resulting from exposure misclassification can be calculated as the normalized difference between the expected value and true value of the logistic regression coefficient using % Attenuation Bias = (β_E - β_T)/ β_T *100 (Whitehead et al. 2012).

We used linear mixed-effects models with random intercepts to identify significant predictors (p<0.1) of U-Cd levels and estimate the amount of variability in measured levels explained by the model while accounting for the correlation among repeat samples collected from the same individual (Peretz et al. 2002). We included creatinine concentration as a predictor in our models with unadjusted U-Cd levels as the dependent variable instead of using creatinine-adjusted U-Cd as the dependent variable because this allows for an evaluation of the relationship between U-Cd and other predictor variables independent of urinary creatinine concentration (Barr et al. 2005). We first identified potential explanatory variables from questionnaire data that were associated with U-Cd concentrations with p<0.2 in either bivariate or linear trend analyses. We then used a backwards stepwise elimination regression to derive final multivariate models including all questionnaire variables that predicted U-Cd concentrations with p<0.1, in addition to the environmental and dietary Cd exposure estimates that were retained in final models regardless of

their p-values. To estimate associations with passive tobacco smoke exposure, we created a separate model restricted to women who never smoked (n=163). We used information on the years since former smokers (n=70) quit smoking by adding three years to data collected in 1997 to estimate the percentage decrease in U-Cd concentrations per year after smoking cessation adjusted for other significant predictors. We performed a 10-fold cross-validation to evaluate the fit of our models by setting aside 10% of the data and rerunning the models (Shao 1993). All analyses were performed using SAS version 9.2 (SAS Institute, Inc., Cary, NC, USA) and STATA version 11 (STATA Corp., College Station, TX, USA).

RESULTS

Table 1 presents the characteristics of the study participants. At enrollment into the CTS, substudy participants had a median of two full-term pregnancies, breastfed for a median duration of three months, and most had never smoked (68%). At the time of urine collection in 2000, substudy participants were on average 55 years of age, had a slightly greater than ideal body mass (median = 25 kg/m²) and an average body surface area of 1.8 m². Traffic density (0 – 427,000 VKT/km²) and industrial Cd emissions (0 – 1,760 kg) at participant residences in 2000 ranged over several orders of magnitude, while interquartile ranges for estimated dietary Cd intake and Cd concentrations in ambient air were 7.9 – 14 μ g/day and 0.09 – 0.28 η g/m³, respectively. The geometric means of U-Cd concentration, creatinine-adjusted U-Cd concentration, and 24-hour U-Cd output from the first 24-hour urine sample were 0.27 μ g/L, 0.38 μ g/g and 0.46 μ g/day, respectively.

The ICC among the 141 participants with repeated 24-hour urine samples was 0.50 for U-Cd concentration and 0.42 for creatinine-adjusted U-Cd concentration, indicating moderate within-

person correlation over time. The ICC was similar whether the time between repeat urine sample collection was 3, 6 or 9 months (ICC=0.51, 0.59 and 0.42 respectively). Based on the overall within- and between-person variance components (0.221 and 0.216, respectively; a variance ratio of 1.0), measurement error resulting from the use of a single 24-hour U-Cd sample to estimate exposure would result in a 50% attenuation bias of the regression coefficient towards the null while the use of two or four U-Cd samples would result in 33% and 20% attenuation bias respectively.

Table 2 presents selected results from analyses of self-reported characteristics and U-Cd levels in the first 24-hour urine sample (n=296). U-Cd levels increased significantly (p<0.1) with both age and cumulative pack-years of smoking, and the relationship was stronger for creatinine-adjusted U-Cd concentrations (p<0.0001). Passive smoking was not related to U-Cd concentrations using either categorical or continuous measures. The geometric mean creatinine-adjusted U-Cd concentration among those with 20 or more pack-years of smoking (0.57 µg/g) was 63% higher than the geometric mean levels among never-smokers (0.35 μ g/g). There was a significant linear trend (p=0.0001) of decreasing U-Cd concentrations as category of average daily alcohol consumption increased from none to <20g/day (one drink) to ≥20g/day (two or more drinks). Increasing parity was also related to lower U-Cd levels (p = 0.02) and was a stronger predictor of creatinine-adjusted U-Cd levels (p = 0.0002). Larger body surface area was associated with lower unadjusted (p = 0.04) and creatinine-adjusted U-Cd (p=0.08), while higher body mass index was associated with lower creatinine-adjusted U-Cd (p=0.08) but not unadjusted U-Cd concentrations (p = 0.14). Duration of breastfeeding (p=0.05) and ever use of hormone replacement therapy (p=0.02) were associated with creatinine-adjusted U-Cd but not unadjusted U-Cd. Never using oral contraceptives was associated with higher unadjusted (p=0.002) and

creatinine-adjusted U-Cd concentrations (p=0.001). There was no relationship between U-Cd or creatinine-adjusted U-Cd levels and estimated dietary Cd intake. Although U-Cd levels increased with categorical estimated exposure to all three environmental sources of Cd, these differences were not significant nor was there a significant linear trend with the exception of industrial cadmium emissions within five kilometers and unadjusted U-Cd concentrations (p=0.06) (Table 3).

Table 4 provides the estimated percentage change in U-Cd concentrations from the final mixedeffects models based on all available 24-hour urine samples with smoking as a predictor variable for all participants (Model 1) and among never-smokers with passive smoking intensity-years as a predictor variable (Model 2). The total variance explained (R^2) was similar at 42 - 44% for both models. The greatest variability in U-Cd concentrations was explained by creatinine concentration (27%) and age (8%). Total pack-years of smoking among all participants and total lifetime intensity of passive smoking among nonsmokers were also positively associated with U-Cd. Each year in age was associated with a 1.4% increase (95% CI: 0.9, 1.9) in U-Cd concentration and each pack year of active smoking was associated with a 1% increase (95% CI: 0.5, 1.6). Among former smokers (n=70 participants and 97 samples), the number of years since smoking stopped was associated (p=0.01) with a 1.5% decrease (95% CI: -2.7, -0.3) per year in U-Cd concentration (data not shown). Increasing parity and alcohol intake were associated with lower U-Cd concentrations. Dietary and environmental estimates of Cd exposure were not significant predictors (p>0.1) of U-Cd concentrations in this population. Body surface area, body mass index, breastfeeding duration, use of hormone replacement therapy and oral contraceptives were dropped from the final multivariate models because they were not significant predictors (p>0.1).

Models with creatinine-adjusted U-Cd or 24-hour U-Cd output as the dependent variable produced parameter estimates similar to those for U-Cd concentration (results not shown). Cross-validation showed that the models were not over fit, with the same independent variables significant in each subset of the data, similar regression coefficients ($\pm 10\%$) and overall adjusted R^2 values (40-46%).

DISCUSSION

In this analysis, we observed only a moderate level of within-person correlation for repeated measures of U-Cd concentrations (unadjusted ICC=0.50; creatinine-adjusted ICC=0.42) from 24-hour samples collected 3-9 months apart, suggesting that a single measurement does not accurately represent lifetime Cd body burden. This result from repeat 24-hour urine samples is within the range of correlations (r=0.4-0.6) observed from the few studies that measured U-Cd in repeat morning void samples (Ikeda et al. 2006; Mason et al. 1998; Yamagami et al. 2008). For an epidemiologic study of the effect of Cd where a single measurement of U-Cd is used to characterize exposure, our observed variance ratio (within to between-person variance) of one for repeated samples leads to exposure misclassification with an estimated attenuation bias of approximately 50%, such that a "true" odds ratio of 2.0 would be reduced to an observed value of 1.4. Creatinine-adjusted U-Cd levels in this study (GM=0.38 μ g/g) were nearly identical to estimates from other studies in the U.S. in women of similar age (GM=0.28 – 0.36 μ g/g) that reported associations between U-Cd and cancer mortality (Adams et al. 2012), breast cancer incidence (McElroy et al. 2006) and cardiovascular mortality (Menke et al. 2009).

In this population of California women without occupational exposure to Cd and a very low prevalence of current smoking (3%), we identified several factors that predicted U-Cd

concentrations. Age and lifetime pack-years of smoking were positively associated with U-Cd, consistent with previous studies (Adams et al. 2011; Ikeda et al. 2005; McElroy et al. 2007a; Richter et al. 2009). The association with age, however, may be due to age-related changes in renal physiology such as lower Cd excretion among older individuals due to reduced tubular reabsorption capacity (Bernard 2004; Vahter et al. 2004; Jarup and Akesson 2009) as well as lower absorption of Cd in older women due to postmenopausal increases in serum ferritin (Milman et al. 1992; Jian et al. 2009). Each pack-year of smoking was associated with a 1% increase in U-Cd concentrations in our study population of women who were 31 -84 years of age, compared with a 2% increase estimated for a population of premenopausal women who were 40 –45 years of age (Adams et al. 2011). Lifetime intensity of passive smoke exposure was also associated with U-Cd among never-smokers in our population. One study observed a significant correlation (r=0.5, p=0.02) between urinary cotinine and U-Cd among 23 children from Sweden (Willers et al. 2005), while a study of 254 women in Wisconsin who were 30 - 69years of age found no association between U-Cd and self-reported recent passive smoke exposure duration or number of locations where women were exposed although no measure of smoke intensity was included (McElroy et al. 2007a).

We observed a weak negative relationship between parity and U-Cd concentration, a finding consistent with a recent study of premenopausal women (Adams et al. 2011). Other studies observed a positive association between U-Cd and parity that the study authors attributed to potential iron deficiency during pregnancy resulting in increased absorption of Cd (Akesson et al. 2002; McElroy et al. 2007b). Average daily alcohol consumption was inversely associated with U-Cd in our study population; this contradicts previous studies that reported no association with any alcohol consumption compared to none (Gil et al. 2011) and ordinal variables similar to

our measure of alcohol consumption (McElroy et al. 2007b; Peters et al. 2010). Consistent with previous studies, body surface area, a measure of muscle mass, was inversely associated with U-Cd in bivariate models (Dhooge et al. 2010; McElroy et al. 2007b; Suwazono et al. 2005); however, neither body surface area nor body mass index, a measure of adiposity, were associated with U-Cd in adjusted models (data not shown).

Studies of populations consuming food contaminated with Cd have observed positive associations between dietary Cd intake and urinary Cd levels (Ikeda et al. 2006; Yamagami et al. 2006) as have several other studies in low-exposure populations (Choudhury et al. 2001; Julin et al. 2011; Shimbo et al. 2000). However, consistent with our findings, other studies of women with low Cd exposure have observed no association between U-Cd levels and either dietary Cd intake (Vahter et al. 1996) or the consumption of specific food items (McElroy et al. 2007b). In a study of non-smoking women, Adams et al. (2011) observed an association between U-Cd and usual consumption of tofu and cooked cereals. The Total Diet Study, from which we obtained our estimates of dietary Cd in the foods that participants reported consuming, did not include Cd levels in tofu or other soy products. However, tofu was not largely consumed in our population based on food frequency questionnaires. While some studies measured Cd in duplicate food samples (Julin et al. 2011; Shimbo et al. 2000; Vahter et al. 1996) and others relied on food frequency questionnaires and Total Diet Study data (Choudhury et al. 2001; and the present study), no clear pattern between methodologies and results was apparent. Variation in Cd absorption related to iron stores is another possible explanation for the mixed findings between dietary and urinary Cd (McElroy et al. 2007; Vahter et al., 1996). In addition, Cd levels in food are dependent on soil levels as evidenced by a study in Japan that showed a two-fold variation in Cd levels in foods grown in different locations (Shimbo et al. 2000).

We did not observe associations between estimated Cd exposures from environmental sources and U-Cd in multivariate models. Our study area, which included urban, suburban, and rural regions of the state, was not known to have Cd contamination and had relatively low Cd emissions compared to areas of California with a greater concentration of industrial sources. Furthermore, the ambient levels of Cd observed in the study area are considered to be low and not thought to be a major source of exposure to the general population (ATSDR 1999). Our GIS-derived exposure estimates were based on residential location in 2000 only, and did not account for time spent in other locations (e.g., workplace), wind direction or meteorology. Nonetheless, the amount of variance in U-Cd concentrations explained by our mixed-effects regression models (R²=42-44%) is similar to that observed in a study (R²=40%) of non-occupationally exposed women that used a single measure of creatinine-adjusted U-Cd (McElroy et al. 2007b).

Limitations of this analysis include the relatively low and limited range of U-Cd levels and low levels of estimated Cd exposure from smoking and environmental sources among our study participants. Our estimates of potential environmental exposure were based on a single year of data and may not accurately reflect historical cumulative exposure through air. Urine samples were stored for about 9 years at -70 °C and were not collected specifically for U-Cd analysis, therefore potential contamination of the urine collection containers that could contribute to the observed within- and between-person variation in measured concentrations could not be ruled out (McElroy et al. 2007b). Other than tobacco smoke, we were unable to identify sources of exposure that were statistically significant predictors of U-Cd levels. We did not measure iron status in our participants which can be an important factor influencing gastrointestinal uptake of Cd (Akesson et al. 2002; Gallagher et al. 2011; Julin et al. 2011; Satarug et al. 2010). We also

did not have any information on renal function such as measures of glomerular filtration that were positively and paradoxically associated with U-Cd in recent studies, suggesting a potential reverse causality (Chaumont et al. 2012; Weaver et al. 2011a; Weaver et al. 2011b). The strengths of the present study include use of 24-hour urine samples (Akerstrom et al. 2012), the ability to evaluate within-person variation in U-Cd levels for about half of the study subjects; comprehensive questionnaire data related to dietary, reproductive and lifestyle factors; and the estimation of Cd exposure from outdoor sources.

CONCLUSIONS

These results suggest that urinary cadmium levels increase with age and exposure to tobacco smoke and that a single measurement of 24-hour urinary cadmium does not accurately reflect medium to long-term (i.e. 6-9 month average) body burden.

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Table 1. Distributions of demographic, lifestyle, and geographic factors and laboratory data from the first 24-hour urine samples provided by participants.

Variable (units)	N	Min	25 th	50 th	75 th	Max	Min	25 th	Mean±SD
Self-reported data									
Age (years)	296	31	47	54	62	84	31	47	55±12
Body mass index	293	16	23	25	29	61	16	23	27 ± 5.9
(kg/m^2)									
Body surface area	293	1.2	1.7	1.8	1.9	2.7	1.2	1.7	1.8 ± 0.2
(m^2)									
Parity (full-term	291	0	0	2	3	6	0	0	1.7±1.4
pregnancies)									
Breastfeeding	290	0	1	3	5	9	0	1	3.2 ± 2.3
(months)									
Smoking (pack-	296	0	0	0	1.0	61	0	0	3.9±10
years)									
Passive smoking ^a	171	0	4.0	21	40	203	0	4.0	30±35
(intensity-years)									
Dietary Cd intake ^b	287	2.5	7.9	11	14	28	2.5	7.9	11±4.2
(µg/day)									
Geographic									
exposure data									
Industrial	296	0	0	0	0.003	1,760	0	0	18±146
emissions ^c (kg)									
Outdoor air (ng/m ³)	296	0.05	0.09	0.15	0.28	0.65	0.05	0.09	0.20 ± 0.13
Traffic density ^d	296	0	0	7,000	33,900	427,000	0	0	27,500±55,200
(VKT/km^2)									
Urinary									
concentrations									2
Unadjusted Cd	296	0.1	0.2	0.3	0.4	2.0	0.1	0.2	0.27 ± 1.9^{e}
$(\mu g/L)$									2
Creatinine (g/L)	296	0.1	0.5	0.7	1.0	2.5	0.1	0.5	0.71 ± 1.6^{e}
Creatinine-adjusted	296	0.1	0.3	0.4	0.5	1.5	0.1	0.3	0.38 ± 1.8^{e}
$\operatorname{Cd}(\mu g/g)$									0.15.1 = 9
Cd output (μg/day)	295	0.1	0.3	0.5	0.7	2.6	0.1	0.3	0.46 ± 1.7^{e}

^aAmong never-smokers responding to a 1997 questionnaire about exposure to second-hand smoke (n=171), calculated by multiplying duration of exposure in years by intensity (1=little smoky, 2=fairly smoky, 3=very smoky). ^bDietary cadmium intake estimated from food frequency questionnaire combined with cadmium levels measured in food items from the Total Diet Study (USFDA 2006). ^cCadmium emissions within five kilometers of residence. ^dVKT = vehicle kilometers traveled; traffic density = VKT per square kilometer within 300 meters of residence. ^eGeometric mean and geometric standard deviation.

Table 2. Selected host and lifestyle characteristics and 24-hour urinary cadmium concentrations from the first sample provided, unadjusted and adjusted for creatinine.

			Una	djusted Cd (p	ug/L)	Creatinine-adjusted Cd (µg/g)			
Characteristic	n	%	GM (μg/L)	Kruskal- Wallis p-value ^a	Linear trend p- value ^b	GM (μg/g)	Kruskal-Wallis p-value ^a	Linear trend p-value ^b	
Age (years)									
31 - 39	30	10%	0.26	0.11	0.01	0.30	< 0.0001	< 0.0001	
40 - 49	63	21%	0.23			0.31			
50 - 59	114	39%	0.27			0.38			
60 - 84	89	30%	0.30			0.45			
Smoking (pack-years)									
0 (never)	207	70%	0.26	0.02	0.01	0.35	0.0002	< 0.0001	
0.1 - 4.9	40	13%	0.25			0.40			
5.0 - 19.9	26	9%	0.27			0.43			
\geq 20	23	8%	0.41			0.57			
Passive smoking									
(intensity-years) ^c									
< 4	42	25%	0.26	0.15	0.70	0.34	0.27	0.53	
4 - 20	44	27%	0.31			0.40			
21 - 40	43	25%	0.24			0.34			
> 40	37	23%	0.28			0.37			
Alcohol (g/day)									
None	96	32%	0.32	0.0005	0.0001	0.44	0.0006	0.0001	
< 20	174	59%	0.25			0.36			
≥ 20	26	9%	0.19			0.29			
Parity (full-term									
pregnancies)									
0	75	26%	0.28	0.65	0.02	0.41	0.39	0.0002	
1 - 2	143	49%	0.27			0.37			
3	43	15%	0.24			0.35			
> 3	30	10%	0.25			0.37			
Total duration of									
breastfeeding (months)									
≤1	81	27%	0.28	0.37	0.17	0.41	0.03	0.05	
$\frac{-}{2} - 3$	81	27%	0.29	,		0.40	****		
4 - 5	87	29%	0.25			0.33			
> 5	47	16%	0.25			0.36			
Oral contraceptive use									
Ever	201	72%	0.25	0.002	-	0.35	0.001	-	
Never	80	28%	0.32	0.002		0.44	0.001		
Hormone replacement		-0,0	0.52			· · · ·			
therapy									
Ever	141	48%	0.27	0.31	_	0.38	0.02	_	
Never	154	52%	0.27	0.51		0.37	0.02		
Body mass index	101	52/0	0.27			0.57			
(kg/m ²)									
< 25.0	145	50%	0.28	0.30	0.14	0.40	0.15	0.08	
25.0 – 29.9	82	28%	0.24	0.50	0.17	0.40	0.13	0.00	
≥ 30.0	66	22%	0.24			0.35			
Body surface area (m ²)	00	22/0	0.20			0.50			
< 1.65	73	25%	0.31	0.04	_	0.42	0.08	_	
< 1.05 ≥ 1.65	223	75%	0.25	0.04	-	0.42	0.00	-	
_ 1.03	443	15/0	0.23			0.50			

^aNon-parametric test using the Kruskal-Wallis one-way analysis of variance by ranks. ^bLinear test for trend from regression model of natural logarithm-transformed concentrations adjusted for age. ^cAmong never-smokers responding to the 1997 questionnaire about exposure to second-hand smoke (n=171), calculated by multiplying duration of exposure in years by intensity (1=little smoky, 2=fairly smoky, 3=very smoky).

Table 3. Dietary and environmental characteristics and 24-hour urinary cadmium concentrations from the first sample provided, unadjusted and adjusted for creatinine.

			Unad	justed Cd (μ	g/L)	Creatinine Adjusted Cd (µg/g)		
Characteristic	n	%	GM (μg/L)	Kruskal- Wallis p-Value ^a	Linear trend p-Value ^b	GM (μg/g)	Kruskal- Wallis p-Value ^a	Linear trend p-Value ^b
Dietary cadmium intake								
(µg/day)								
< 7.9	71	25%	0.28	0.72	0.30	0.38	0.82	0.98
7.9 - 10.6	72	25%	0.27			0.37		
10.6 - 13.7	73	25%	0.27			0.37		
> 13.7	71	25%	0.25			0.38		
Estimated outdoor cadmium concentration (ng/m ³)								
< 0.1	94	32%	0.25	0.37	0.29	0.36	0.34	0.31
0.1 - 0.3	142	48%	0.27			0.38		
> 0.3	60	20%	0.29			0.39		
Traffic density (vehicle								
kilometers traveled per square								
kilometer)								
0	101	34%	0.27	0.78	0.82	0.37	0.64	0.54
1 - 7,000	48	16%	0.27			0.36		
7,001 - 70,000	118	40%	0.26			0.38		
> 70,000	29	10%	0.29			0.41		
Industrial cadmium emissions								
(kg within 5 kilometer)								
0	203	69%	0.26	0.23	0.06	0.37	0.57	0.23
0.001 - 20	52	18%	0.27			0.38		
> 20	41	14%	0.30			0.40		

^aNon-parametric test using the Kruskal-Wallis one-way analysis of variance by ranks.

^bLinear test for trend from regression model of natural logarithm-transformed concentrations adjusted for age.

Table 4. Estimated adjusted percentage change and 95% confidence interval (95% CI) in 24-hour urinary cadmium concentration (μ g/L) associated with potential predictors.

Variable (categories if applicable)	Model 1 (all participants) % change ^a (95% CI)	p- value	\mathbb{R}^2	Model 2 (never-smokers) % change ^a (95% CI)	p- value	R ²
Number of (Samples/subjects)	412/285			233/161		
Creatinine (per 0.1 g/L)	15 (12, 19)	< 0.001	0.27	14 (11, 19)	< 0.001	0.27
Age ^b (per year)	1.4 (0.9, 1.9)	<0.001	0.35	1.1 (0.40, 1.8)	0.003	0.35
Smoking ^c (per lifetime pack-year)	1.1 (0.5, 1.6)	0.001	0.37	Not Included		
Passive smoking ^d (per lifetime intensity-year)	Not Included			0.2 (0.0, 0.4)	0.05	0.37
Total full-term pregnancies ^e (per pregnancy)	-4.6 (-8.6, -0.5)	0.004	0.39	-5.2 (-10.1, 0.0)	0.05	0.39
Alcohol intake ^b $(0, <20, \ge 20 \text{ g/day})$	-16 (-23, -7.4)	< 0.001	0.40	-16 (-24, -4.9)	0.01	0.40
Cadmium in air ^b (per 0.1 ng/m ³)	-1.4 (-7.9, 5.5)	0.77	0.40	1.0 (-7.4, 11)	0.28	0.41
Industrial emissions ^b - within 5 km (per 10-fold change in kg)	2.5 (-1.9, 7.1)	0.27	0.41	1.5 (-4.4, 7.7)	0.63	0.41
Traffic density ^b - within 300m (per 10-fold change in VKT ^f /km ²)	-1.9 (-4.9, 1.2)	0.22	0.41	-2.2 (-5.8, 1.5)	0.24	0.42
Dietary cadmium intake ^g (per µg/day)	-0.1 (-1.4, 1.4)	0.64	0.42	-0.9 (-2.5, 0.7)	0.16	0.44

^aPercent change = $[\exp(\beta)-1]*100$. ^bQuestionnaire or residence from first urine collection in 2000.

^cQuestionnaire from follow-up in 1997. ^dAmong never-smokers responding to the 1997 questionnaire about exposure to second-hand smoke (n=171), calculated by multiplying duration of exposure in years by intensity (1=little smoky, 2=fairly smoky, 3=very smoky). ^eTotal live and still births from questionnaire at cohort enrollment in 1995-96. ^fVKT= vehicle kilometers traveled per day. ^gDietary cadmium intake estimated from food frequency questionnaire from 2000 combined with cadmium levels measured in food items from the Total Diet Study (USFDA 2006).